

A Hierarchical Bayesian Model for Ocean Properties Reconstructions

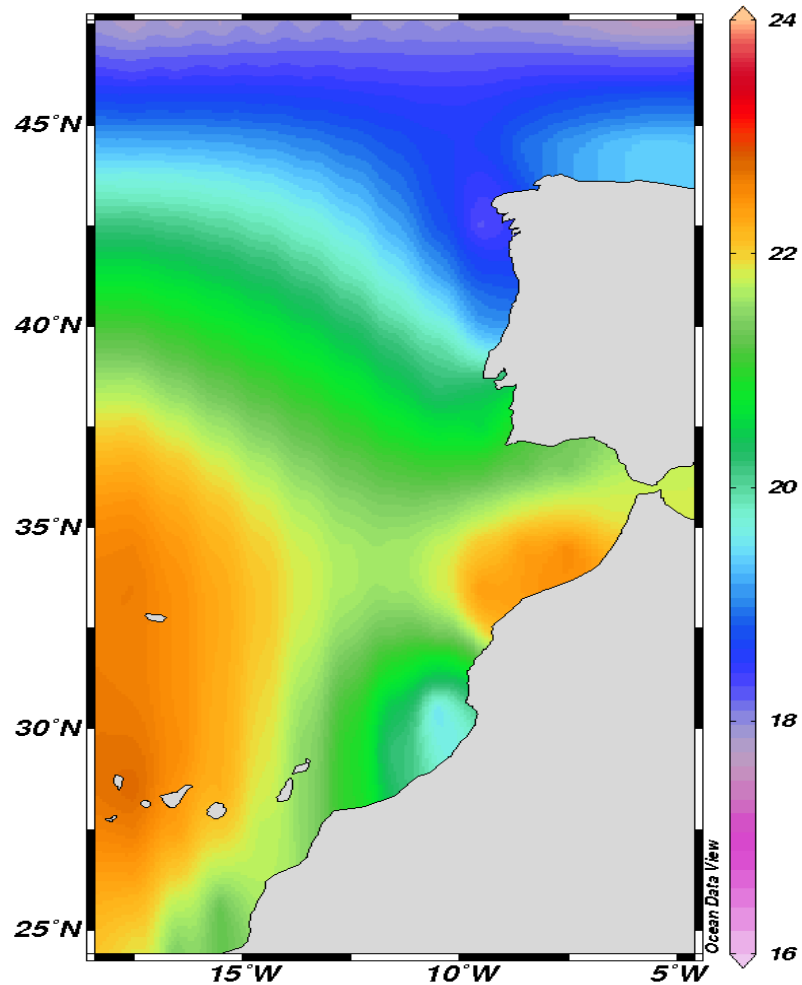
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In collaboration with Ricardo Lemos, NOAA-NMFS, Pacific Grove,
CA.

MOTIVATING ISSUE

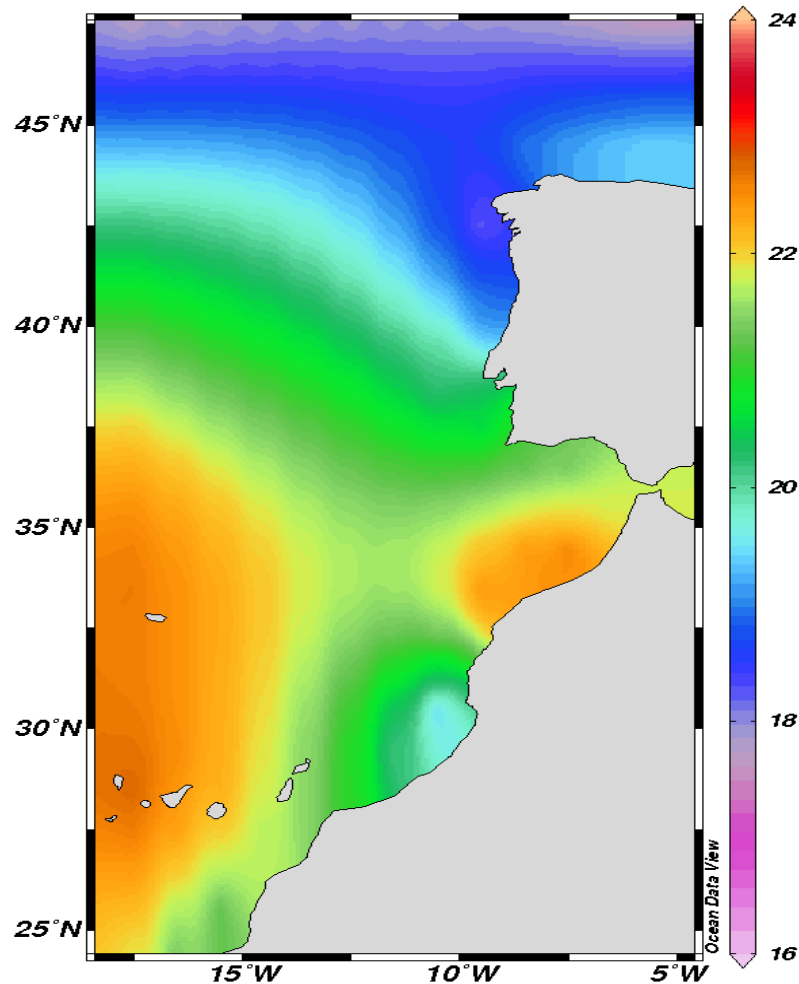
AUGUST AVERAGE SST FROM WOA98



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- There is no upwelling off the Iberian peninsula and a very weak one in Western Africa.

GOALS

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- The method must be useful in large geographical domains and long time frames.
- Observational errors should be accounted for.

BAYESIAN HIERARCHICAL MODELS

We propose a statistical model based on the Bayesian hierarchical structure. This consists of three layers:

- **Observation Equation:** relationship between observations and the true state of nature.
- **Process equation:** probabilistic description of the space and time variability of the state vector.
- **Prior distributions:** prior knowledge on the parameters that define the previous two layers.

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- All estimation uncertainty is propagated through the different levels of the model and fully accounted for.
- These models lend themselves to the use of Monte Carlo methods. This provides for easy probabilistic inference for quantities derived from the model parameters.

SPATIAL VARIABILITY

A non-homogeneous spatially-varying process can be written as

$$X(\mathbf{s}) = \sum_{i=1}^m b(\mathbf{s} - \mathbf{s}_i^*; \boldsymbol{\omega}(\mathbf{s})) \gamma_i, \quad \boldsymbol{\gamma} \sim N(0, \mathbf{K})$$

where, $\mathbf{s}_1^*, \dots, \mathbf{s}_m^*$ is a grid, and

$$b(\mathbf{s} - \mathbf{j}; \boldsymbol{\omega}) \equiv \begin{cases} (1 - \|\mathbf{s} - \mathbf{j}\|_{\boldsymbol{\Sigma}}^2)^{\omega_1} & \text{if } \|\mathbf{s} - \mathbf{j}\|_{\boldsymbol{\Sigma}} < 1 \\ 0 & \text{otherwise.} \end{cases}$$

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The distance is given as

$$\|\mathbf{s} - \mathbf{j}\|_{\boldsymbol{\Sigma}} \equiv \sqrt{((x_s - x_j), (y_s - y_j)) \boldsymbol{\Sigma}^{-1} ((x_s - x_j), (y_s - y_j))^T}.$$

SPATIAL VARIABILITY

The ellipsoidal shape is controlled by the parameters in

$$\Sigma^{-1} \equiv \begin{pmatrix} \Psi_1 + \Psi_2 \cos 2\pi\omega_4 & \Psi_2 \sin 2\pi\omega_4 \\ \Psi_2 \sin 2\pi\omega_4 & \Psi_1 - \Psi_2 \cos 2\pi\omega_4 \end{pmatrix}$$

$$\Psi = \frac{1}{2} \left(\frac{1}{a^2} + \frac{1}{A^2}, \frac{1}{a^2} - \frac{1}{A^2} \right)$$

$$a = L + \omega_2(U - L), \quad A = a + \omega_3(U - a), \quad \omega_2, \omega_3 \in (0, 1)$$

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The spatial variation of $\boldsymbol{\omega}$ is obtained, with a normalized b , as

$$\boldsymbol{\omega}(\mathbf{s}) = \sum_{i=1}^m b(\mathbf{s} - \mathbf{s}_i^*; \mathbf{u}) \boldsymbol{\rho}(\mathbf{s}_i^*) \quad \mathbf{u} = (2, 1, 0, 0)$$

with appropriate uniform priors on each $\rho_k(\mathbf{s}_i^*)$, $k = 1, \dots, 4$.

SST RECONSTRUCTION

The SST observation $x_{i,m,y}(\mathbf{s})$ corresponding to data set $i = 1, \dots, 4$ (OSD, CTD, XBT, MBT), in month m , year y and location \mathbf{s} follows

$$x_{i,m,y}(\mathbf{s}) = \theta_{m,y}(\mathbf{s}) + \varepsilon_{i,m,y}(\mathbf{s}), \quad \varepsilon_{i,m,y}(\mathbf{s}) \sim N(0, \tau_i^2).$$

Different τ_i allow for different observational error variances depending on the instrument.

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Anomaly:

$$\Delta_{m,y}(\mathbf{s}) = \theta_{m,y}(\mathbf{s}) - \bar{\Xi}_m(\mathbf{s}).$$

SST RECONSTRUCTION

The true SST is

$$\theta_{m,y}(\mathbf{s}) \sim N \left(\sum_{\mathbf{j}} b(\mathbf{s} - \mathbf{j}; \mathbf{\Lambda}(\mathbf{s})) \gamma(\mathbf{j}), \Phi(\mathbf{s})^2 \right)$$

where

$$\gamma(\mathbf{j}) = \alpha(\mathbf{j}) + \beta_t(\mathbf{j}) \mathbf{w}_t^T + \eta(\mathbf{j})(t - 180),$$

\mathbf{j} denotes a points in a 4° resolution grid, and

$$\Phi(\mathbf{s})^2 = \sum_{\mathbf{j}} b(\mathbf{s} - \mathbf{j}; \mathbf{\Omega}(\mathbf{s})) \exp(\sigma(\mathbf{j})).$$

Here $t = m + 12(y - 1961)$ denotes time in months since December 1960. Note that $\theta_{m,y}(\mathbf{s})$ is continuous in space.

Lack of stationarity in time is handled by letting

$$\boldsymbol{\beta}_t \sim N(\boldsymbol{\beta}_{t-1}, \mathbf{W}_t)$$

and

$$\mathbf{w}_t = \left(\sin\left(\frac{2\pi t}{12}\right), \cos\left(\frac{2\pi t}{12}\right), \sin\left(\frac{2\pi t}{6}\right), \cos\left(\frac{2\pi t}{6}\right) \right).$$

Where \mathbf{W}_t is modeled using a space-dependent discount factor.

IMPLEMENTATION

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We use reasonably vague inverse gamma priors for all variance parameters. Posterior inference shows that the data do provide information about those parameters.

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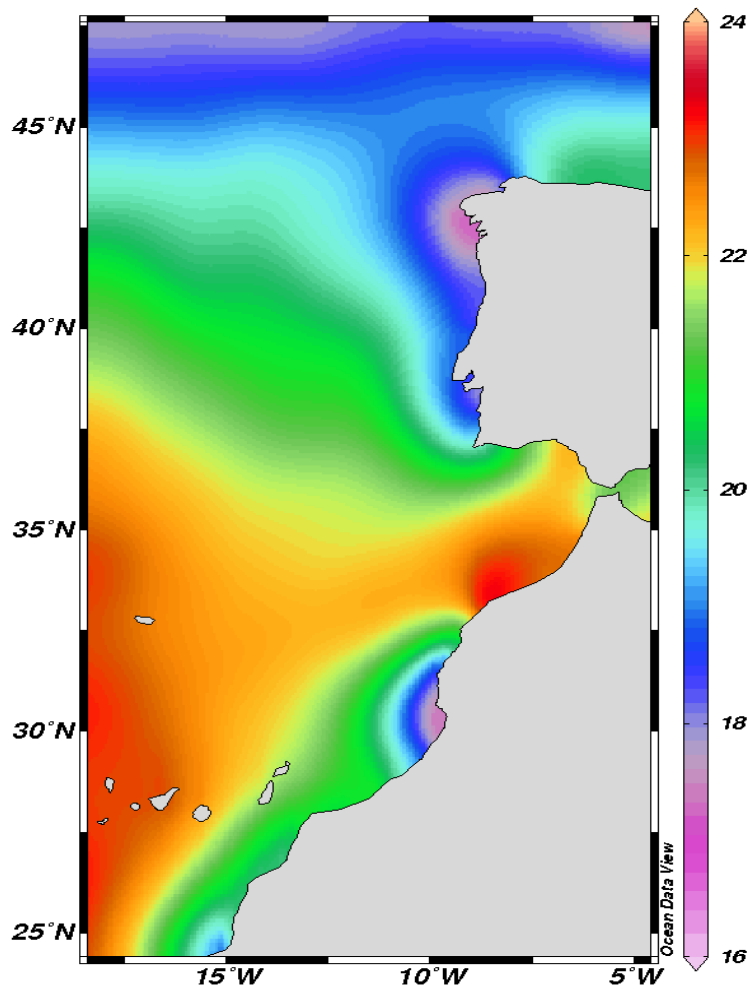
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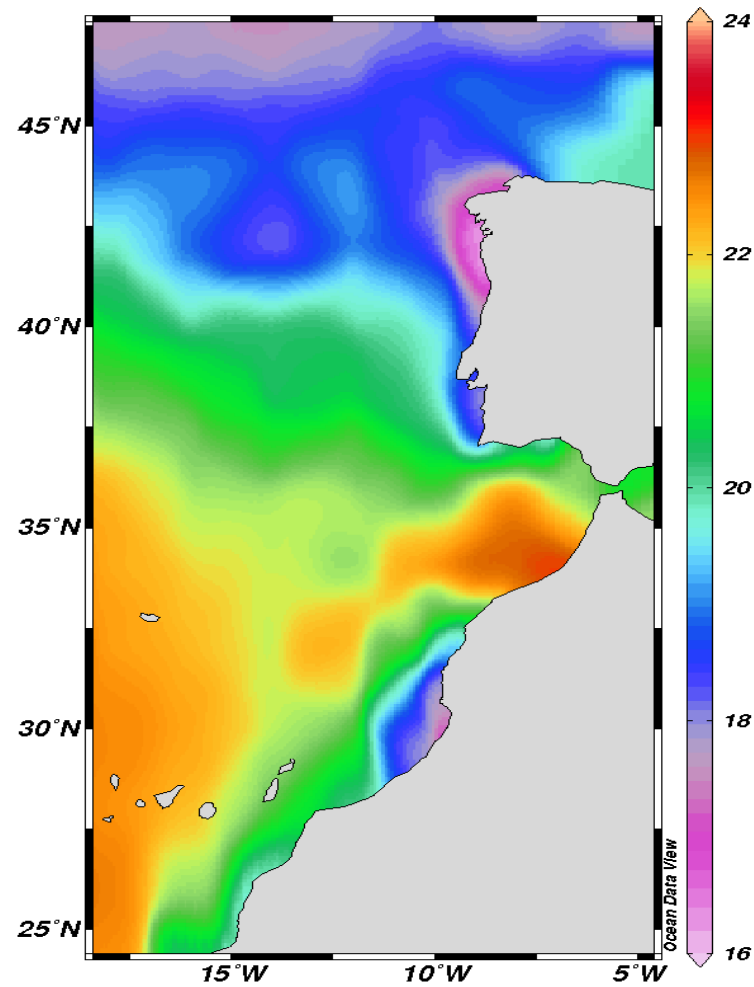
We run an MCMC. To determine convergence we use the diagnostics available in BOA to set the burn-in (1,200 iterations), the thinning (1/3) and the sample size (6,000 from the thinned chain). We also performed two separate runs, a warm start configuration (30°) and a cold one (15°).

RESULTS: AUGUST CLIMATOLOGY

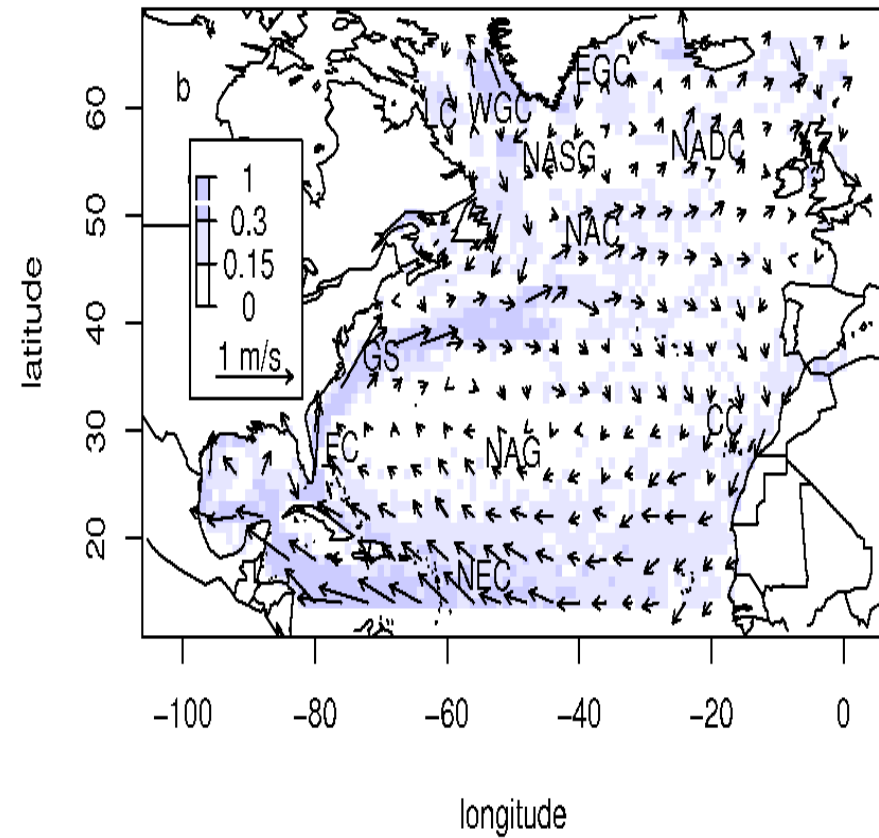
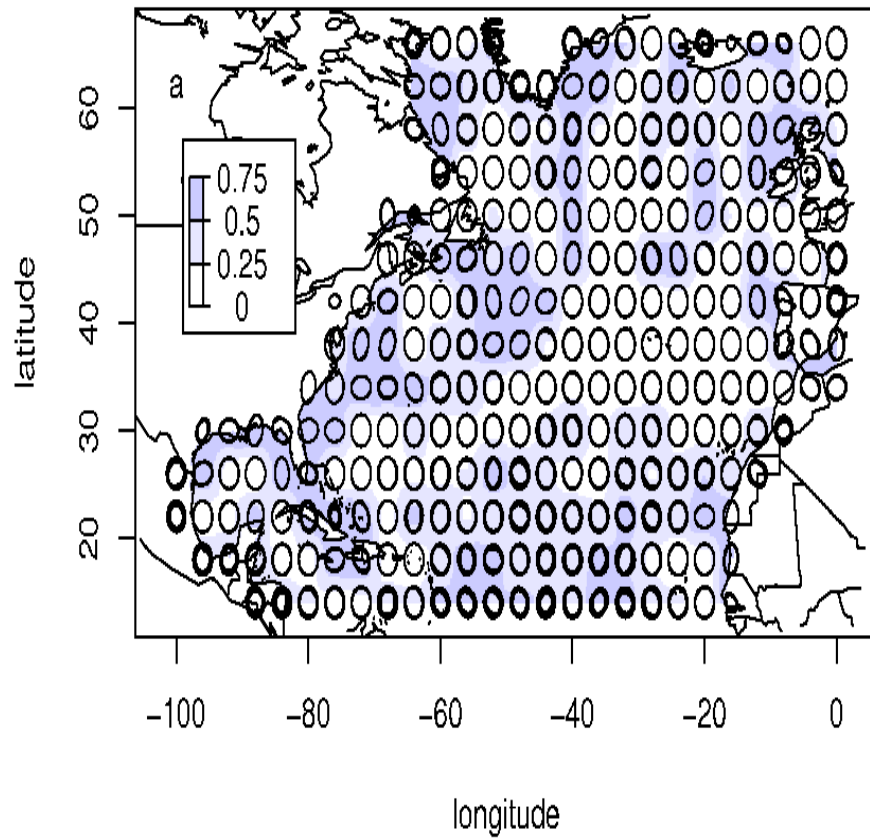
AUGUST SST FROM WOA01



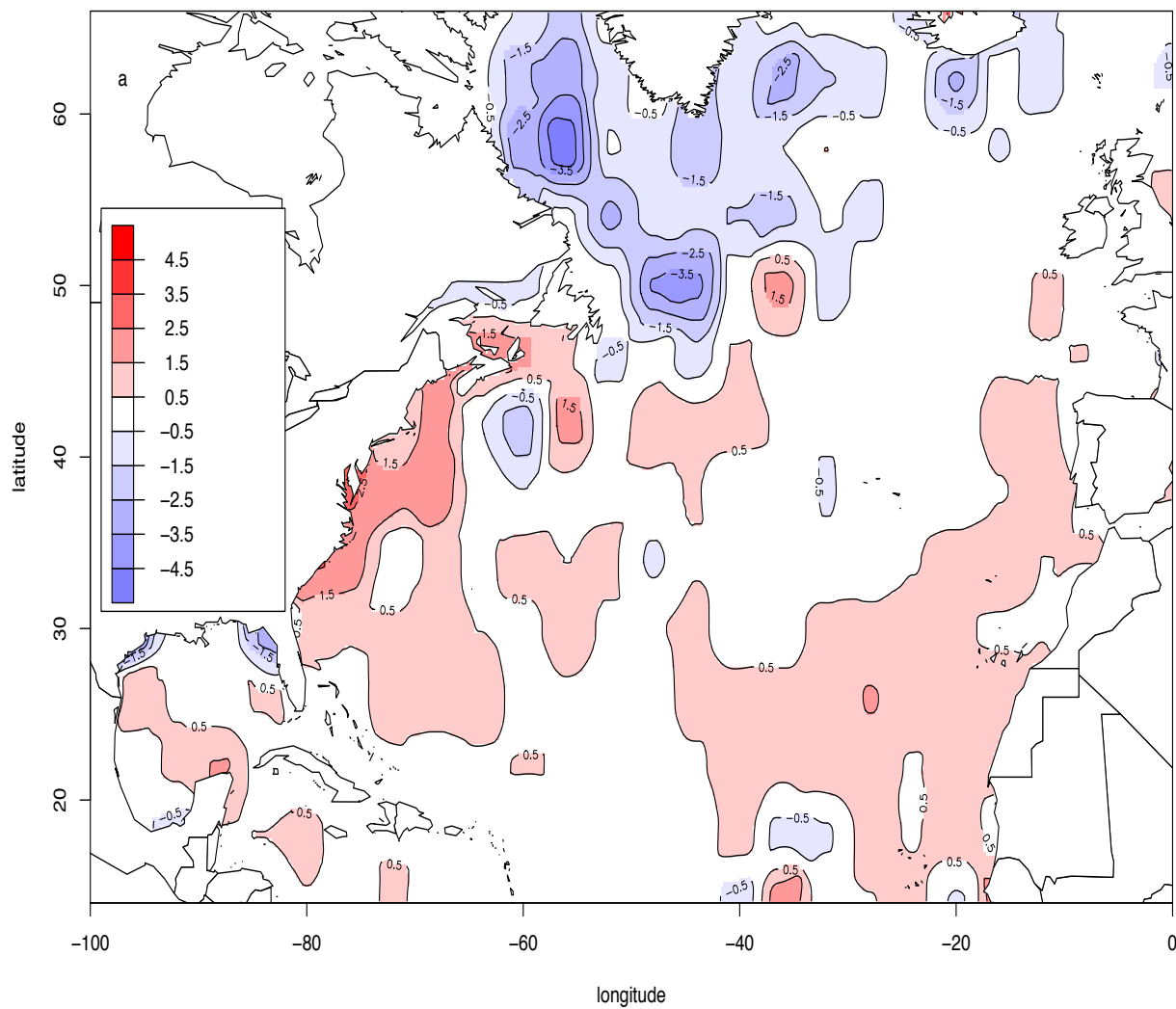
AUGUST SST FROM LS09



RESULTS: KERNELS VS CURRENTS



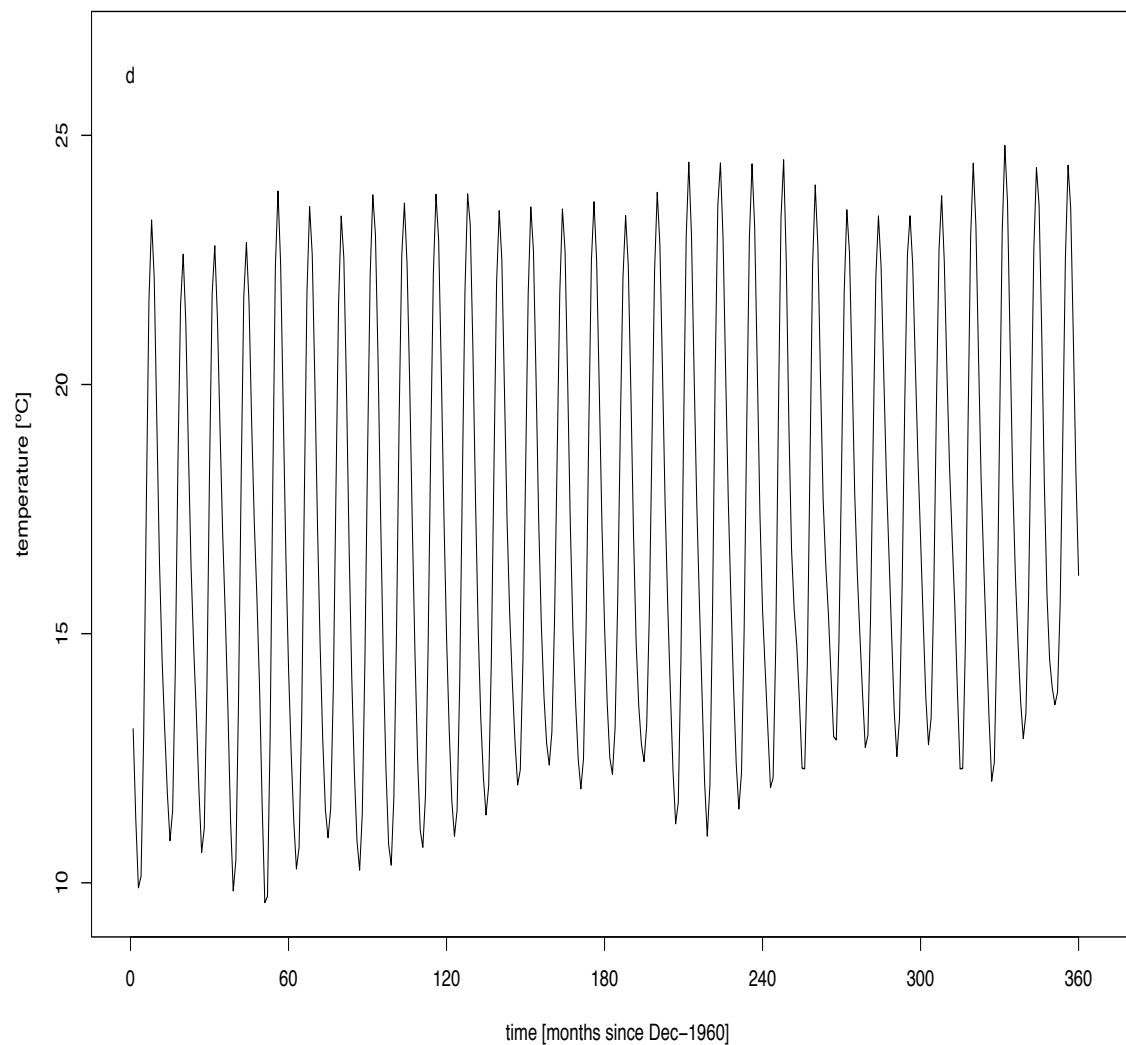
RESULTS: TRENDS



Trends
($^{\circ}\text{C}/30$
years)

RESULTS: TRENDS

Posterior mean
for monthly SST
at s_2 ($^{\circ}\text{C}$).



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- The column of water needs to satisfy a density stability constraint. That is, density must increase with depth. Density is related to temperature and salinity via the equation of state.
- The previous two points imply that we need to consider a hierarchical structure for our statistical model that incorporates some physical constraints.

AN OCEAN MODEL

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- Temperature T and salinity S
- Pressure p and density ρ .

Based on observations that we denote as \mathbf{Y} we want to build a statistical model to explore the posterior distribution

$$\pi(\mathbf{u}, T, S, p, \rho | \mathbf{Y})$$

We write

$$\pi(\mathbf{u}, T, S, p, \rho | \mathbf{Y}) = \pi(T, S | \mathbf{u}, \mathbf{Y}) \pi(\mathbf{u} | \rho, p, \mathbf{Y}) \pi(p | \rho, \mathbf{Y}) \pi(\rho | \mathbf{Y})$$

For the remaining of the talk I will focus on the model for density, $\pi(\rho | \mathbf{Y})$.

OBSERVATIONAL MODEL

We let $\mathbf{Y}_t = (Y_t(\mathbf{s}_1), \dots, Y_t(\mathbf{s}_{n_t}))^T$, be the vector of densities at time t for all locations where records are available at that time.

Then

$$\mathbf{Y}_t \sim N(\mathbf{F}^o \boldsymbol{\rho}_t, \text{diag}(\mathbf{u}_t))$$

where \mathbf{F}^o is an incidence matrix and $\boldsymbol{\rho}_t$ is a vector of stacked true, unobserved, density profiles on a $.25^\circ$ resolution grid.

PROCESS MODEL

Following the same framework of the SST model, we let

$$\boldsymbol{\rho}_t \sim N(\mathbf{F}(\boldsymbol{\mu}_{m(t)} + \mathbf{W}\boldsymbol{\gamma}_t), \tau\mathbf{I})$$

$m(t)$ is the month corresponding to time t . Thus, there are 12 vectors $\boldsymbol{\mu}_t$ that determine the regular seasonal cycle.

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The elements of $\boldsymbol{\mu}_t$ are sampled from a multivariate normal truncated to impose that its components are in increasing order. \mathbf{F} is obtained from Bezièr kernels centered on a 1° resolution grid. Whence, the product $\mathbf{F}\boldsymbol{\mu}_t$ has its components in increasing order.

PROCESS MODEL

\mathbf{W} is a matrix of spatial factors and $\boldsymbol{\gamma}_t$ is a low dimensional vector capturing time-varying trends and low frequency cycles. Thus

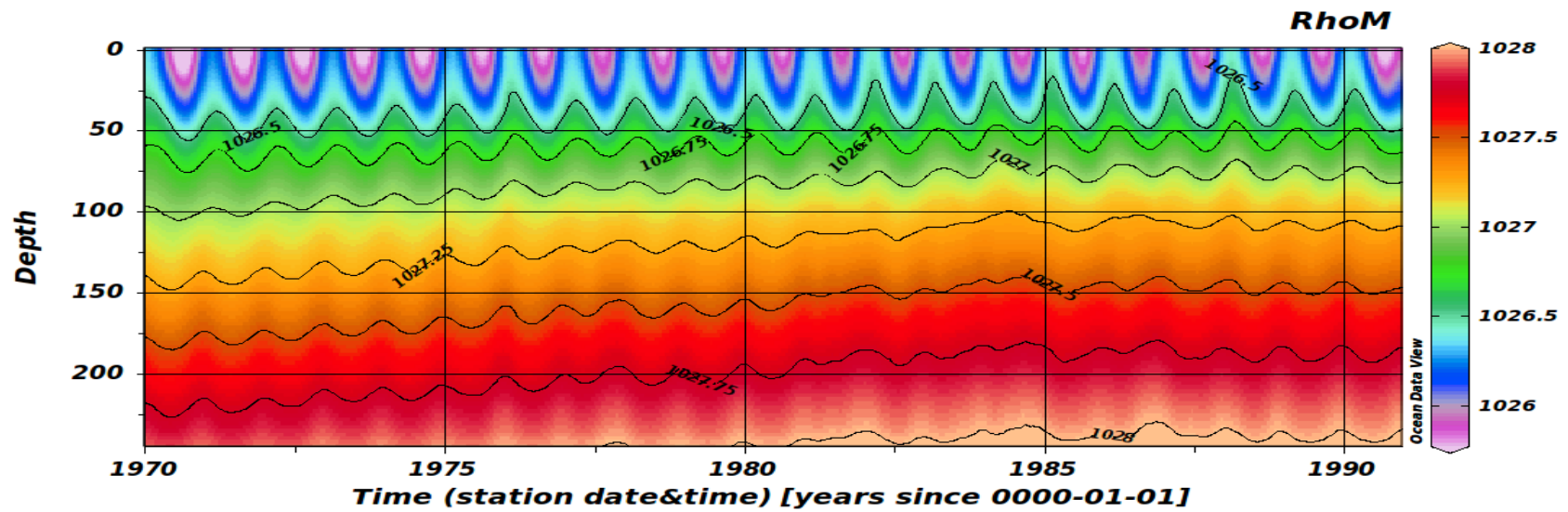
$$\boldsymbol{\gamma}_t \sim N(\mathbf{G}\boldsymbol{\gamma}_{t-1}, \mathbf{H})$$

for a rotation matrix \mathbf{G} and an appropriately chosen matrix \mathbf{H} .

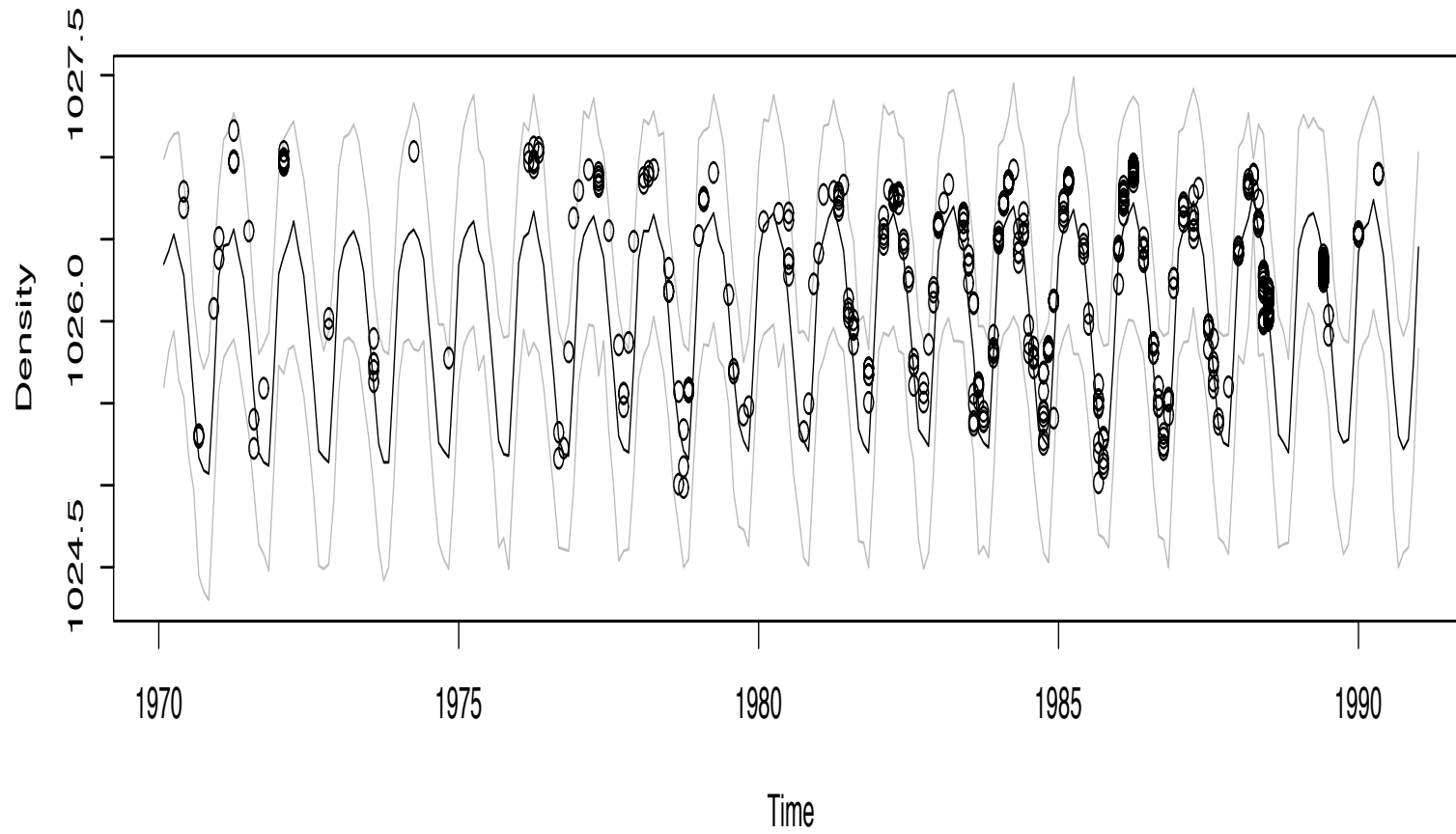
PRELIMINARY RESULTS

We considered a $2^\circ \times 2^\circ$ grid with 33 layers in the vertical centered at 39N 12W (off Lisbon). We used NODC's OSD and CTD data from 1970 to 1991

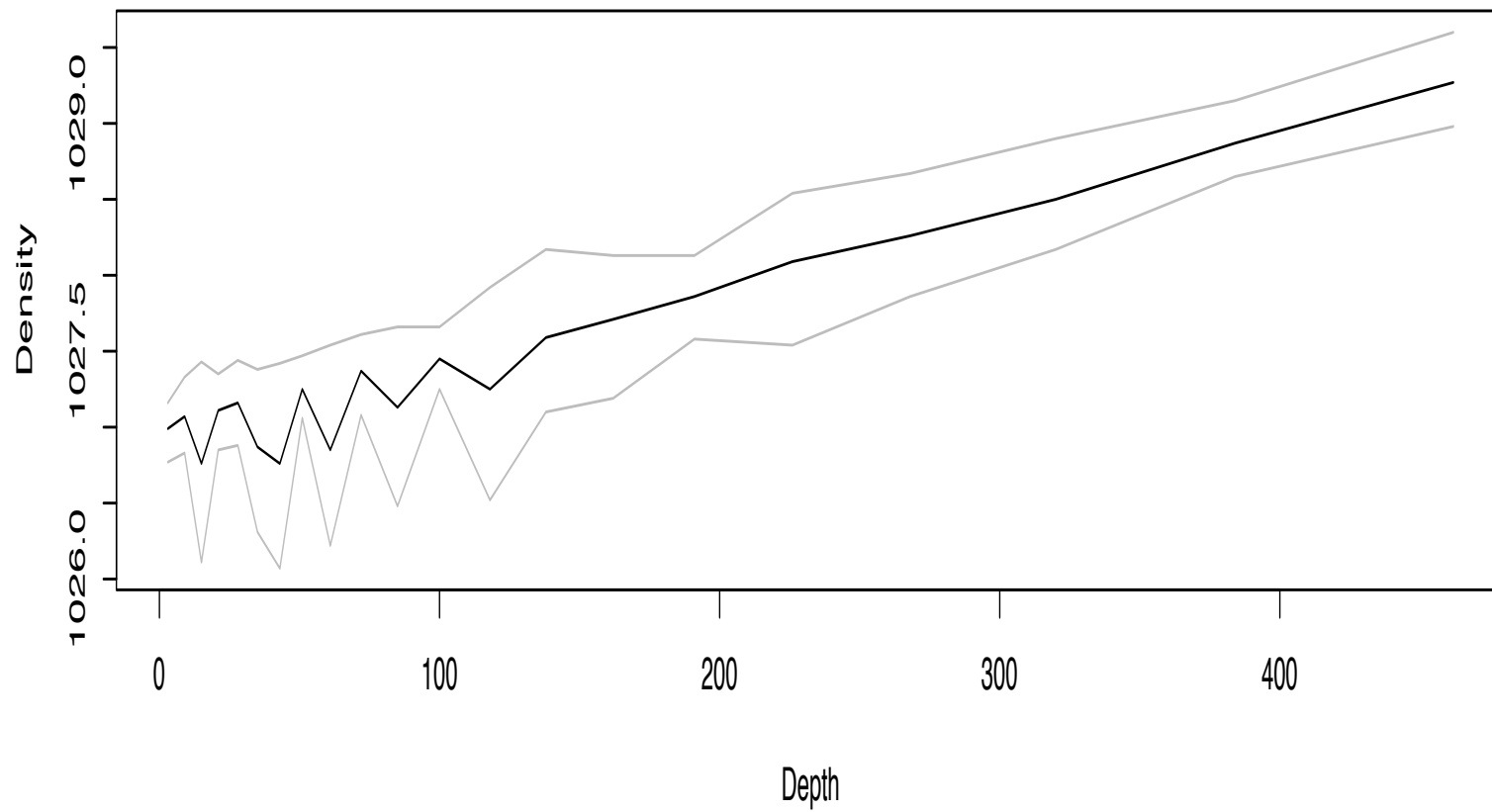
DENSITY PROFILE



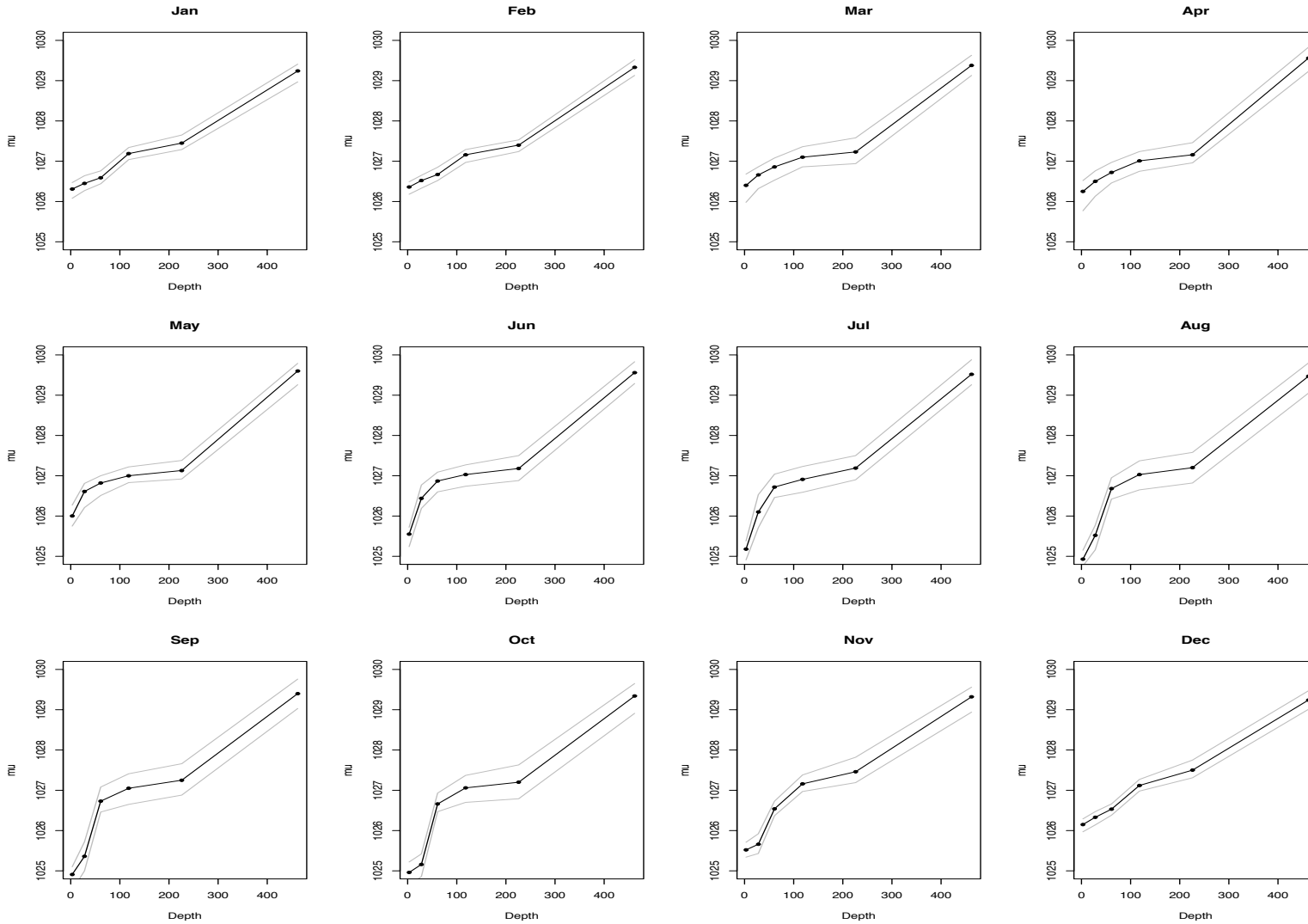
SURFACE DENSITY



SINGLE LOCATION PROFILE



MEAN DENSITY PROFILE



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- Our model includes observational errors and realistic descriptions of the latent processes governing the evolution of ocean variables.
- The model is able to provide probabilistic assessments of the variabilities included in the estimated quantities. All estimation variabilities are accounted for in the final product.
- Our model is able to handle large data sets. By using kernels with compact support and making use of the structure of the CDLM we are able to parallelize the estimation algorithms.

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- Reference: Ricardo T. Lemos, Bruno Sansó (2006)
“Spatio-temporal Variability of Ocean Temperature in the Portugal Current System”. *Journal of Geophysical Research Oceans*, 111, C04010, doi:10.1029/2005JC003051.

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- Reference: Ricardo T. Lemos, Bruno Sansó (2009) “A Spatio-Temporal Model for Mean, Anomaly and Trend Fields of North Atlantic Sea Surface Temperature (with discussion)”. *Journal of the American Statistical Association*, 104, pp. 5–25.
Winner of the 2010 Mitchell Prize to the best applied Bayesian paper!